# **B34.UC2**

# Numerical Computation and Statistics in Engineering

Unit 5: Hypothesis Testing

## **Hypothesis Testing**

Statistical tests consist of the following elements:

- $\bullet$  a *null hypothesis*  $H_0$  about one or more population parameters;
- an alternative hypothesis  $H_1$  (or  $H_a$ ) that replaces  $H_0$  if the test does not support  $H_0$ ;
- the test statistics;
- acceptance and rejection regions indicating the values of the statistics that will lead to acceptance or rejection of  $H_0$ .

The term null hypothesis stems from the fact that we often test for 'something being equal to 0', for example  $\mu-4=0$  (i.e., the population mean equals 4), or  $\mu_1-\mu_2=0$  (i.e., the two populations have the same mean).



There are two types of errors we can make when testing a hypothesis:

		$H_0$ true	$H_0$ false
Decision:	Reject $H_0$	Type I error	Correct decision
	Accept $H_0$	Correct decision	Type II error

Type I errors (rejecting  $H_0$  while it is true) are usually denoted by the symbol  $\alpha$ , type II errors (accepting  $H_0$  while it is false) are denoted by the symbol  $\beta$ .

**Example.** A car retailer believes that more than 20% of his customers are willing to spend extra money for upgrading the stereo equipment of their new car. Before ordering new equipment the retailer wants to ask 10 of his customers whether they would buy the more expensive equipment.



Here we pick

$$H_0: p = 0.2.$$

$$H_a: p > 0.2.$$

(We do not believe in p < 0.2.) The random variable x for this test is the number of people indicating that they would buy better stereo equipment for their cars. If p = 0.2 we expect  $10 \cdot 0.2 = 2$  people to be in favor of the better product, thus, rejecting  $H_0$  if  $x \ge 4$  seems reasonable. For the type II error we find

$$\alpha = P(\text{reject } H_0 \text{ while it is true})$$

$$= P(p = 0.2 \text{ and } x \ge 4)$$

$$= 1 - P(p = 0.2 \text{ and } x \le 3)$$

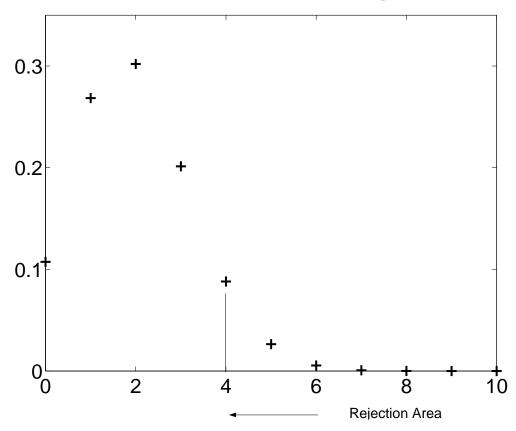
$$= 1 - \sum_{x=0}^{3} {10 \choose x} p^x (1-p)^{10-x}$$

$$\approx 0.121.$$



Questioning the customers the retailer finds that 4 out of 10 people are in favor of the better product, thus  $H_0$  is rejected.

### Binomial Distribution, n = 10, p = 0.2



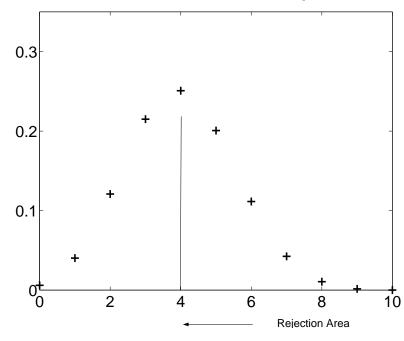


Suppose that the true parameter value is p=0.4. Then

$$\beta = P(\text{accept } H_0 \text{ while } p = 0.4) = P(x \le 3 \text{ while } p = 0.4)$$

$$= \sum_{x=0}^{3} {10 \choose x} 0.4^x (1 - 0.4)^{10-x} \approx 0.3823.$$

#### Binomial Distribution, n = 10, p = 0.4





# **Using the Normal Distribution**

In practice we can often use the normal distribution to find acceptance and rejection intervals. Suppose we want to test

$$H_0: \theta = \theta_0 \qquad H_1: \theta \neq \theta_0$$

where  $\theta$  is a parameter of a population (probability, mean, etc.).  $\theta_0$  is the value that we think  $\theta$  has. We assume that the estimator  $\hat{\theta}$  that we get from the sample has normal distribution with mean  $\theta_0$  and standard deviation  $\sigma_{\hat{\theta}}$ . Then

statistics 
$$z = \frac{\hat{\theta} - \theta_0}{\sigma_{\hat{\theta}}}$$

has a standard normal distribution. If the rejection region is

$$z < -z_{\alpha/2}, \qquad z_{\alpha/2} < z$$

then the type I error is  $\alpha$ .



# **Using the Normal Distribution**

Indeed,

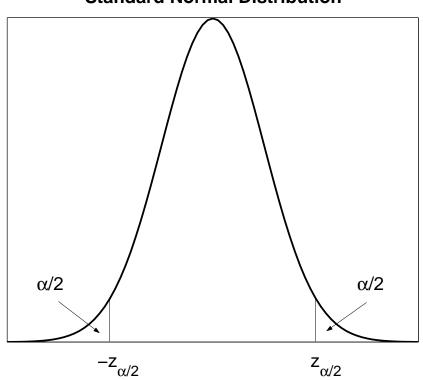
type I error

 $= P(H_0 \text{ holds but is rejected})$ 

$$= P(z \le -z_{\alpha/2} \text{ or } z_{\alpha/2} \le z)$$

 $= \alpha$ .

#### **Standard Normal Distribution**



Such a test is called a two-tailed test.



# **Using the Normal Distribution**

For a *one-tailed test* the data are

- $H_0$ :  $\theta = \theta_0$ ;
- $H_1$ :  $\theta > \theta_0 \ (\theta < \theta_0)$ ;
- Statistics:  $z = \frac{\hat{\theta} \theta_0}{\sigma_{\hat{\theta}}}$ ;
- Rejection region:  $z > z_{\alpha}$   $(z < -z_{\alpha})$ ;
- Type I error:  $\alpha$ .

In practice, we are often given  $\alpha$  in advance specifying the type I error probability that we are willing to accept, and we use this to find the acceptance and rejection interval.



One-tailed test	Two-tailed test	
$H_0$ : $\mu = \mu_0$	$H_0$ : $\mu = \mu_0$	
$H_1: \mu > \mu_0$	$H_1 \colon \mu \neq \mu_0$	
(or $H_1$ : $\mu < \mu_0$ )		
$z = \frac{\bar{x} - \mu_0}{\sigma_{\bar{x}}}$	$z = rac{ar{x} - \mu_0}{\sigma_{ar{x}}}$	
Rejection region:	Rejection region:	
$z>z_{lpha}$ (or $z<-z_{lpha}$ )	$z < -z_{lpha/2}$ or $z_{lpha/2} < z$	

If the sample size is small (n < 30) or if  $\sigma_{\bar{x}}$  has to be estimated by  $s/\sqrt{n}$  then the normal distribution is replaced by the t-distribution with n-1 degrees of freedom.



**Example.** We go back to the machines making wires with diameter approximately 1mm. Data taken from two machines showed the following values:

I: 1.0429 1.0627 0.9203 0.9280 1.0286

 $0.9800 \quad 1.0345 \quad 1.0408 \quad 1.0356 \quad 1.0645$ 

II: 1.0001 1.0157 0.9439 0.9794 0.9753

 $0.9319 \quad 0.9877 \quad 0.9483 \quad 1.0225 \quad 0.9558$ 

with  $n_1=n_2=10$ ,  $\bar{x}_1=1.0138$ ,  $s_1=0.0526$ ,  $\bar{x}_2=0.9761$ , and  $s_2=0.0309$ .

For

$$t_i = \frac{\bar{x}_i - 1.0}{s_i / \sqrt{10}}$$

we find  $t_1 = 0.7877$  and  $t_2 = -2.4459$ . With  $\alpha = 0.05$  we consider the following tests:



- $H_0$ :  $\bar{x}_1 = 1.0$ ,
- $H_1$ :  $\bar{x}_1 < 1.0$ .

Since  $t_1 \not> t_{9,0.05} = 1.8331$  the hypothesis  $H_0$  is accepted.

- $H_0$ :  $\bar{x}_1 = 1.0$ ,
- $H_1: \bar{x}_1 \neq 1.0.$

Since  $t_{9,0.025}=2.2622$  and  $-2.2622 < t_1 < 2.2622$  the hypothesis is accepted.

In both cases the decision is 'correct'; the data was random data with  $\mu=1.0$  and  $\sigma=0.05$ .



For the second sample we consider

- $H_0$ :  $\bar{x}_2 = 1.0$ ,
- $H_1$ :  $\bar{x}_2 < 1.0$ .

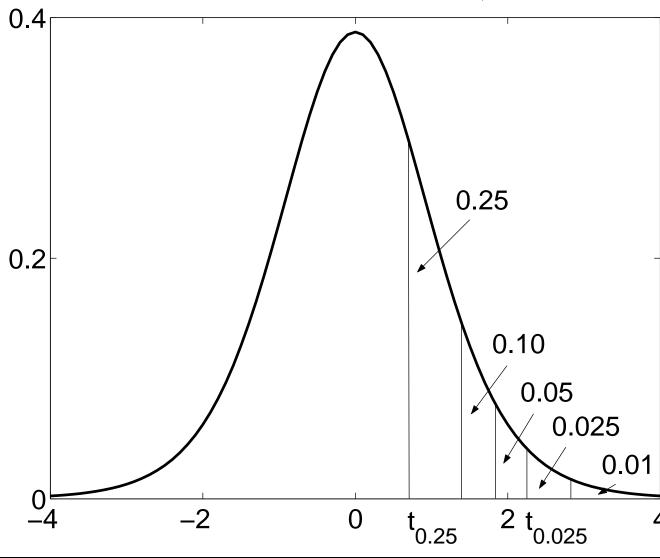
Since  $t_2 = -2.4459 < t_{9,0.05} = -1.833$  the hypothesis is rejected. For the two-tailed test

- $H_0$ :  $\bar{x}_2 = 1.0$ ,
- $H_1: \bar{x}_2 \neq 1.0$

we see that  $t_2 < -t_{9,0.025} = -2.262$ , and we reject  $H_0$  again.









#### One-tailed test

 $H_0: \mu_1 - \mu_2 = d$ 

 $H_1: \mu_1 - \mu_2 < d$ 

(or  $H_1: \mu_1 - \mu_2 > d$ )

$$z = \frac{(\bar{x}_1 - \bar{x}_2) - d}{\sigma_{\bar{x}_1 - \bar{x}_2}} = \frac{(\bar{x}_1 - \bar{x}_2) - d}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

Rejection region:

$$z > z_{\alpha} \text{ (or } z < z_{\alpha})$$

#### Two tailed test

 $H_0: \mu_1 - \mu_2 = d$ 

 $H_1: \mu_1 - \mu_2 \neq d$ 

$$z = \frac{(\bar{x}_1 - \bar{x}_2) - d}{\sigma_{\bar{x}_1 - \bar{x}_2}} = \frac{(\bar{x}_1 - \bar{x}_2) - d}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

Rejection region:

$$z < -z_{\alpha/2}$$
 or  $z_{\alpha/2} < z$ 

If the sample sizes are small and  $\sigma_1$  and  $\sigma_2$  are unknown then the tdistribution with  $n_1 + n_2 - 2$  degrees of freedom replaces the normal distribution, with

$$t = \frac{(\bar{x}_1 - \bar{x}_2) - d}{\sqrt{s_P^2(\frac{1}{n_1} + \frac{1}{n_2})}} \quad \text{where} \quad s_P^2 = \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2},$$

provided that the two unknown variances are equal.

**Example.** The response times of two hard drives are tested. The values found are

Disk 1
 Disk 2

 
$$n_1 = 15$$
 $n_2 = 13$ 
 $\bar{x}_1 = 16$ 
 $\bar{x}_2 = 13$ 
 $s_1 = 5$ 
 $s_2 = 4$ 

What can be said about the difference between the mean response times?



Here

$$H_0: (\mu_1 - \mu_2) = 0, \qquad H_1: (\mu_1 - \mu_2) \neq 0.$$

We calculate

$$s_P^2 = \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2}$$
$$= \frac{14 \cdot 5^2 + 12 \cdot 4^2}{15 + 13 - 2}$$
$$= 20.8462$$

and

$$t = \frac{(\bar{x}_1 - \bar{x}_2) - d}{\sqrt{s_P^2(\frac{1}{n_1} + \frac{1}{n_2})}}$$

$$= \frac{16 - 13}{\sqrt{20.8462(\frac{1}{15} + \frac{1}{13})}}$$

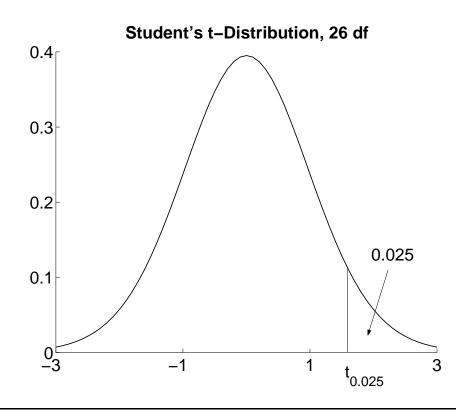
$$= 1.7340.$$



From the tables we know that for  $n_1+n_2-2=26$  and  $\alpha=0.1$  (for example) that

$$t_{0.05,26} = 1.706$$
.

Since  $t_{0.05,26} < t$  the hypothesis is rejected.





**Example.** A study is done in the effectiveness of certain exercises to help weight loss. The following data are collected (in kg):

before	after	before	after
106	99	86	83
90	87	78	77
86	86	92	92
107	104	83	82
91	90	101	100
97	96	90	88
80	81	22	116
91	91	73	71

For the random variables before b and after a we find  $\bar{b}=92.0625$ ,



 $s_b = 12.3584$ ,  $\bar{a} = 90.0625$ , and  $s_a = 11.0843$ . Our hypotheses are

- $\bullet \ H_0 \colon \bar{b} \bar{a} = 0,$
- $H_1: \bar{b} \bar{a} > 0$ ,

and we will test at a 5% level of significance.

Then  $t_{30,0.05} = 1.699$ , and we reject  $H_0$  if  $t \ge 1.699$ . To calculate further,

$$s_P^2 = \frac{(n_b - 1)s_b^2 + (n_a - 1)s_a^2}{n_b + n_a - 2}$$

$$= \frac{15}{30(s_b^2 + s_a^2)}$$

$$= \frac{1}{2}(12.3584^2 + 11.0843^2)$$

$$= 137.79562.$$



For t we find

$$t = \frac{(\bar{b} - \bar{a}) - 0}{\sqrt{s_P^2(\frac{1}{n_b} + \frac{1}{n_a})}}$$

$$= \frac{92.0625 - 90.0625}{\sqrt{137.79562 \cdot \frac{1}{5}}}$$

$$= 0.38098.$$

The hypothesis is accepted, and there is evidence that the exercises help reducing weight.

# **Tests Concerning the Variance**

Given a random sample from a normal population we will test the null hypothesis  $\sigma^2 = \sigma_0^2$  against the alternatives  $\sigma^2 \neq \sigma_0^2$  or  $\sigma^2 < \sigma_0^2$  ( $\sigma^2 > \sigma_0^2$ ). The random variable

$$\chi^2 = \frac{(n-1)s^2}{\sigma_0^2}$$

has a  $\chi^2$  distribution with n-1 degrees of freedom. For a two-tailed test the null hypothesis is rejected if

$$\chi^2 \le \chi^2_{1-\alpha/2,n-1}$$
 or  $\chi^2 \ge \chi^2_{\alpha/2,n-1}$ .

For a one-tailed test and the alternative hypothesis  $\sigma^2 < \sigma_0^2$  we reject  $H_0$  if  $\chi^2 \leq \chi^2_{1-\alpha,n-1}$ .

**Example.** Thickness of a semi-conductor part (in  $10^{-5}$ m) is crucial in a production process. The machine manufacturing these semi-conductors needs to be readjusted if  $\sigma^2 \leq 0.36$ .

If in a sample of 20 measurements we find  $s^2=0.74$ , what can be said at



# **Tests Concerning the Variance**

a  $\alpha = 0.05$  level of significance?

We assume that thickness is normally distributed. Then

- $H_0$ :  $\sigma^2 = 0.36$ ,
- $H_1: \sigma^2 > 0.36$ .

We reject the null hypothesis if  $\chi^2 \geq \chi^2_{0.05,19}=30.144$ . With  $s^2=0.74$ ,  $\sigma^2_0=0.36$  and n=20 we find

$$\chi^2 = \frac{(n-1)s^2}{\sigma_0^2} = \frac{19 \cdot 0.74}{0.36} = 34.944,$$

and the machine needs to be readjusted.

Note that for n=20,  $\sigma_0^2=0.36$ , and  $\alpha=0.05$  the machine needs readjustment if for a sample of 20 measurements the sample variation  $s^2$  is greater or equal than 0.571.



# **Testing for Proportions**

**Example.** Suppose 5 out of 20 transistors are faulty. We test the hypothesis

- $H_0$ : p = 0.5,
- $H_1: p \neq 0.5$ ,

at the 0.05 level of significance.

Instead of determining the rejection and acceptance interval we will find the smallest  $\alpha$  which will reject  $H_0$  (note for the calculation that the binomial distribution is symmetric):

$$\alpha/2 = P(x \le 5)$$

$$= \sum_{x=0}^{5} {20 \choose x} 0.5^{x} 0.5^{20-x}$$

$$= 0.0207,$$

so that  $\alpha = 0.0414$ . Since  $\alpha < 0.05$  we will reject  $H_0$ .



## **Testing for Proportions**

If in a binomial test the size n is large we can use the normal distribution (with or without continuity correction) as an approximation for the random variable x. Then we get

$$H_0 \colon p = p_0$$
 
$$H_1 \colon p \neq p_0$$
 
$$z = \frac{\bar{x} - np_0}{\sqrt{np_0(1 - p_0)}}$$
 or  $z = \frac{(\bar{x} \pm \frac{1}{2}) - np_0}{\sqrt{np_o(1 - p_0)}}$ 

Rejection region:

$$z < -z_{\alpha/2}$$
 or  $z_{\alpha/2} < z$ 

(If we use the correction factor we use a minus when x exceeds  $np_0$ , and a plus when x is less than  $np_0$ .)



# **Testing for Proportions**

For the on-tailed test

$$H_0: p = p_0$$

$$H_1: p > p_0$$

we use the same statistics z as above with rejection interval  $z \geq z_{\alpha}$ .

**Example.** Suppose  $p_0 = 0.2$  and we test

- $H_0$ : p = 0.2,
- $H_1: p < 0.2$ ,

at the 0.01 level of significance. Then, using  $z_{0.01}=2.33$  we have the rejection region  $z\leq -2.33$ . If the test data are n=200, x=22, then

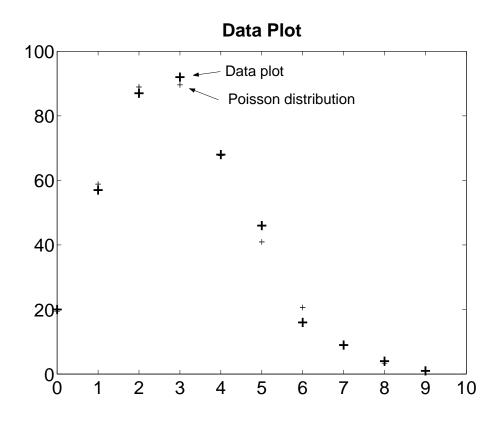
$$z = \frac{\bar{x} - np_0}{\sqrt{np_0(1-p_0)}} = \frac{22 - 200 \cdot 0.2}{\sqrt{200 \cdot 0.2 \cdot 0.8}} \approx -3.18,$$

and we reject the null hypothesis.



Goodness to fit tests are applied to test whether a set of data may be looked upon as a random sample from a population having a given distribution.

Suppose we have data from a Poisson distribution with  $\lambda=3$  (see next slide), which gives the following frequency diagram:





	Frequency		Poisson	Expected freq. $e_i$
x	$f_{i}$		$\lambda = 3$	$(\lambda = 3.0225)$
0	20	0.0500	0.0498	19.4717
1	57	0.1425	0.1494	58.8534
2	87	0.2175	0.2240	88.9421
3	92	0.2300	0.2240	89.6092
4	68	0.1700	0.1680	67.7110
5	46	0.1150	0.1008	40.9313
6	16	0.0400	0.0504	20.6191
7	9	0.0225	0.0216	8.9030
8	4	0.0100	0.0081	3.3637
9	1	0.0025	0.0027	1.1296



For the expected frequency we first estimated  $\lambda$  using the third column as

$$\hat{\lambda} = 3.0225.$$

The random variable

$$\sum_{i=0}^{m} \frac{(f_i - e_i)^2}{e_i}$$

with m the number of different data (here 10) has a  $\chi^2$  distribution with m-t-1 degrees of freedom, where t is the number of parameters estimated from the data (here 1).



With the data above we want to test at a 0.05 level of significance whether the data are from a random variable having Poisson distribution. We set

- $H_0$ : The data are from a Poisson random variable.
- $H_1$ : The data are *not* from a Poisson random variable.

We reject  $H_0$  if

$$\chi_{\alpha,m-t-1}^2 \le \chi^2 = \sum_{i=0}^m \frac{(f_i - e_i)^2}{e_i}$$

Here m=10, t=1, and  $\chi^2_{0.05,10-1-1}=15.507$ . With our data,  $\chi^2=1.9789$ , and  $H_0$  is accepted.



### **Summary**

- Statistical tests often consist of a null hypothesis, and an alternative hypothesis. A type I error is made when the null hypothesis is true, but rejected. A type II error is made when the null hypothesis is false, but is accepted.
- We distinguish one-tailed and two-tailed tests.
- Statistical tests are based on sampling and confidence intervals. We thus use the normal distribution, the Student's *t*-distribution, and the chi-square distribution in standard tests.
- Goodness-To-Fit tests use the chi-square distribution to test whether a set of data fits a given distribution.

